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GENERALIZED CORRELATIONS IN THE SINGULAR CASE. (U)

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BY

10 ASHIS SEN/GUPTA

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OFFICE OF NAVAL RESEARCH

THEODORE W. ANDERSON, PROJECT DIRECTOR

DEPARTMENT OF STATISTICS
STANFORD UNIVERSITY
STANFORD, CALIFORNIA



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ASHIS SEN GUPTA
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GENERALIZED CORRELATIONS IN THE SINGULAR CASE

Ashis Sen Gupta

1. Introduction. When the covariance matrix is singular, the usual expressions for the multiple, partial, canonical [see, e.g., Rao (1973)] and some generalized canonical correlations [for a review, see Sen Gupta (1979)] need to be revised. Tucker, et al. (1972), Khatri (1976) and Rao (1979), (1980) have provided formulae for some of these correlation coefficients in the general case by using g-inverses. A review of their results on multiple, partial and canonical correlations is given first. Next, it is shown that there exists a general representation which covers several generalized canonical correlations and as special cases the multiple, partial and canonical correlations, too. Then a general theory is formulated which deals with the singular case for the representation. Previous results on multiple, partial and canonical correlations follow as special cases of this theory. Further, appropriate formulae are also provided through this formulation for various generalized canonical correlations in the singular case. Finally, the numbers of various critical generalized correlations are derived for the general case.

2. Multiple, partial and canonical correlations in the singular case. Let $R = (R_1: \dots : R_p)$ be the correlation matrix of p variables. Further, let $R^- = (r^{ij}) = (T_1: \dots : T_p)$ be any g-inverse of R . Define, $RR^- = Q = (Q_1: \dots : Q_p)$. Let

I_p have the unit vector e_i , as its i -th column, $i = 1, \dots, p$.

Result 1. The squared multiple correlation of x_1 on x_2, \dots, x_p is

$$\begin{aligned} R_{1 \cdot (2 \dots p)}^2 &= 1 \text{ if } Q_1 \neq e_1 \\ &= 1 - (r^{11})^{-1} \text{ if } Q_1 = e_1 . \end{aligned}$$

Result 2. The partial correlation between x_1 and x_2 eliminating x_3, x_4, \dots, x_p is

$$\begin{aligned} r_{12 \cdot (34 \dots p)} &= 0 \text{ if } Q_1 \neq e_1 \text{ and } Q_2 = e_2 \text{ or if } Q_1 = e_1 \text{ and } Q_2 \neq e_2 \\ &= 1 \text{ if } Q_1 \neq e_1 \text{ and } Q_2 \neq e_2 \\ &= -r^{12}/(r^{11}r^{22})^{1/2} \text{ if } Q_1 = e_1 \text{ and } Q_2 = e_2 . \end{aligned}$$

Let x_1 and x_2 be two sets of variables with the joint dispersion matrix Σ , partitioned accordingly.

Result 3. The squared canonical correlations are the non-zero roots of the determinantal equation

$$|\Sigma_{11}^- \Sigma_{12} \Sigma_{22}^- \Sigma_{21} - \rho^2 I| = 0 \text{ where } \Sigma_{11}^- \text{ and } \Sigma_{22}^- \text{ are any g-inverses of } \Sigma_{22} \text{ and } \Sigma_{11} \text{ respectively.}$$

For proofs and further discussions on the results see Rao [(1979), (1980).]

3. Generalized canonical correlations in the singular case. Canonical correlations have been generalized in various ways. Formulae in the general case will be provided here for those obtained by extending the concepts of tests of independence for two sets of

variates--giving rise to partial, part and bipartial canonical correlations [see Timm (1975) pp. 352-353] and g_1 -, g_2 - bipartial canonical correlations [see Lee (1978)] and some association measures [see McKeon (1965), pp. 16-19]. Various other authors [Horst (1961); Edgerton and Kolbe (1936); Wilks (1938); Lord (1958)] arrived at the same solution as that of McKeon for the particular case of a single variable per set. Appropriate formula will also be provided for the new generalized canonical correlation arising out of the concept of minimum generalized variance [proposed by Anderson (1958) Problem 5, pp. 305-306 and derived by the author {see SenGupta (1979)} under constraint of equi-correlation structure of the generalized canonical variables].

Let $X = (X_1, \dots, X_k)$, $X_i : p_i x_1, p_1 + \dots + p_k = p$, $Disp(X) = k \Sigma$, $Cov(X_i, X_j) = \Sigma_{ij}$ and non-zero ρ s be the generalized canonical correlations. Starting with the defining equations it can be easily seen that for all the above cases, the generalized canonical correlations are obtained from the eigen values of $k \Sigma^*$ in the metric of $k \Sigma_d^*$, i.e. from the solutions of

$$(3.1) \quad |k \Sigma^* - \rho^* k \Sigma_d^*| = 0$$

where $\rho^* = 1 + (k - 1)\rho$ and $k \Sigma_d^*$ is a diagonal super matrix with elements Σ_{ii}^* $i = 1, \dots, k$. In the notation of Lee, $k \Sigma^*$, with $k = 2$, is the covariance matrix of the residual vectors $(\tilde{e}_{1.34}, \tilde{e}_{2.35})$ and $(e_{1.34}, e_{2.35})$ for the g_1 - and

g_2 -bipartial canonical correlations, respectively. In the notation of Timm, $k^{\Sigma^*} = \Sigma_{\cdot 3}$, $\Sigma_{1(2 \cdot 3)}$ and $\Sigma_{(1 \cdot 4)(2 \cdot 3)}$ with $k = 2$ for partial, part and bipartial canonical correlations respectively. Also for McKeon's and the new generalized canonical correlations, $k^{\Sigma^*} = k^{\Sigma}$.

Theorem. The generalized canonical correlations, for the methods quoted above, are given by $\rho = (\rho^* - 1)/(k - 1)$ where ρ^* 's are the non-zero roots of $|k^{\Sigma^*} k_d^{\Sigma^{*-}} - \rho^* I| = 0$, $k_d^{\Sigma^{*-}}$ being any g-inverse of $k_d^{\Sigma^*}$.

Proof. First note the representation (3.1). Consider next the following Lemmas.

Lemma 1. Let A be a hermitian matrix of order n and rank s , and B be non-negative definite matrix of order n and rank r such that $S(A) \subset S(B)$ [where $S(M)$ represents the vector space spanned by the column vectors of M].

Then

(i) There exists a matrix L of order $n \times r$ such that $L^* A L = \Lambda$, $L^* B L = I_r$, where Λ is a diagonal matrix with s non-zero elements, some of which may be repeated and I_r is the identity matrix of order r .

(ii) The non-zero elements of Λ are the same as the roots of the determinantal equation, $|AB^- - \lambda I| = 0$ with repetitions allowed, for any g-inverse B^- of B .

Proof of Lemma 1. See Lemma 3, Rao (1979).

Lemma 2. $S(k^{\Sigma^*}) \subset S(k_d^{\Sigma^*})$.

Proof of Lemma 2. Note that $S(\Sigma_{ij}^*) \subset S(\Sigma_{ii}^*)$ for all the k^{Σ^*} considered above. This follows immediately from the

result [see Proposition 3.31, pp. 3.15-3.16 of Eaton

that, if $A = \begin{pmatrix} A_{11} & A_{12} \\ A_{21} & A_{22} \end{pmatrix} \geq 0$ then $N(A_{22}) \subset N(A_{12})$ and

and $N(A_{11}) \subset N(A_{21})$ where $N(M)$ is the null space of M .

Then, there exist matrices B_{ij} such that $B_{ij} \Sigma_{jj}^* = \Sigma_{ij}^*$, $i, j = 1, \dots, k$.
Hence, there exist matrices C_i such that $(C_1 \dots C_k) \Sigma_d^* = C_k \Sigma_d^* = \Sigma^*$
which proves Lemma 2.

Coupling Lemma 2 with Lemma 1 proves the Theorem.

Note: For $k = 2$, if $p_1 = 1$, $p_2 > 1$ and if $p_1 > 1$, $p_2 > 1$
then we have the cases of multiple and canonical correlations
respectively. Further, with $k = 2$, consideration
of residual variables leads to partial correlation. Thus
the above Theorem unifies the Results 1 through 3 and also
considers simple (and not squared) multiple, partial and
canonical correlations.

4. Numbers of critical generalized correlations.

Let A and B be two hermitian matrices and B be non-negative
definite. If λ is a constant and v a vector such that
 $Av = \lambda Bv$, $Bv \neq 0$, then λ is called a proper eigen value and
 v a proper eigen vector of A with respect to B . In the
context of Lemma 1, the elements of Λ are called the proper
eigen values and the corresponding columns of L , the proper
eigen vectors of A with respect to B . For the generalized
correlations, we consider from (3.1) only the proper eigen

values of k^{Σ^*} with respect to $k_d^{\Sigma^*}$. Also note that for $k \geq 2$, 1 and $-1/(k-1)$ are the maximum and minimum possible values, respectively, for the generalized correlations. Let $k_{od}^{\Sigma^*}$ be the super off-diagonal matrix such that $k^{\Sigma^*} = k_d^{\Sigma^*} + k_{od}^{\Sigma^*}$. Also let $R(M)$ denote the rank of the matrix M .

Lemma 3. The numbers of zero, unit and $-1/(k-1)$ -valued generalized correlations are given by $r = R(k_{od}^{\Sigma^*})$, $r = R[k_{od}^{\Sigma^*} - (k-1)k_d^{\Sigma^*}]$ and $r = R(k_d^{\Sigma^*})$ respectively, where $r = R(k_d^{\Sigma^*})$.

Proof: The proof follows by rewriting (3.1) as

$$|k^{\Sigma^*} - \lambda k_d^{\Sigma^*}| = 0 \text{ where}$$

$(k^{\Sigma^*}, \lambda) = [k_{od}^{\Sigma^*}, (k-1)\rho]$, $[k_{od}^{\Sigma^*} - (k-1)k_d^{\Sigma^*}, (k-1)(\rho-1)]$ and $[k_d^{\Sigma^*}, (k-1)\rho + 1]$ for the zero, unit and $-1/(k-1)$ -valued generalized correlations respectively and noting the one-one relationship between λ and ρ .

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